

Effectiveness of prophylactic intranasal photodynamic disinfection therapy and chlorhexidine gluconate body wipes for surgical site infection prophylaxis in adult spine surgery

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Background: Current measures to prevent spinal surgical site infection (SSI) lack compliance and lead to antimicrobial resistance. We aimed to examine the effectiveness of bundled preoperative intranasal photodynamic disinfection therapy (nPDT) and chlorhexidine gluconate (CHG) body wipes in the prophylaxis of spine SSIs in adults, as well as determine our institutional savings attributable to the use of this strategy and identify adverse events reported with nPDT-CHG.

Methods: We performed a 14-year prospective observational interrupted time-series study in adult (age > 18 yr) patients undergoing emergent or elective spine surgery with 3 time-specific cohorts: before rollout of our institution's nPDT-CHG program (2006–2010), during rollout (2011–2014) and after rollout (2015–2019). We used unadjusted bivariate analysis to test for temporal changes across patient and surgical variables, and segmented regression to estimate the effect of nPDT-CHG on the annual SSI incidence rates per period. We used 2 models to estimate the cost of nPDT-CHG to prevent 1 additional SSI per year and the annual cumulative cost savings through SSI prevention.

Results: Over the study period, 13 493 patients (mean 964 per year) underwent elective or emergent spine surgery. From 2006 to 2019, the mean age, mean Charlson Comorbidity Index (CCI) score and mean Spine Surgical Invasiveness Index (SSII) score increased from 48.4 to 58.1 years, from 1.7 to 2.6, and from 15.4 to 20.5, respectively ($p < 0.001$). Unadjusted analysis confirmed a significant decrease in the annual number (74.6 to 26.8) and incidence (7.98% to 2.67%) of SSIs with nPDT-CHG ($p < 0.001$). After adjustment for mean age, mean CCI score and mean SSII score, segmented regression showed an absolute reduction in the annual SSI incidence rate of 3.36% per year ($p < 0.001$). The estimated annual cost to prevent 1 additional SSI per year was about \$1350–\$1650, and the estimated annual cumulative cost savings were \$2 484 856–\$2 495 016. No adverse events were reported with nPDT-CHG.

Conclusion: Preoperative nPDT-CHG administration is an effective prophylactic strategy for spinal SSIs, with significant cost savings. Given its rapid action, minimal risk of antimicrobial resistance, broad-spectrum activity and high compliance rate, preoperative nPDT-CHG decolonization should be the standard of care for all patients undergoing emergent or elective spine surgery.

Contexte : L'application des mesures de prévention des infections des plaies opératoires (IPO) du rachis laisse actuellement à désirer et pourrait contribuer à la résistance aux antibiotiques. Nous avons voulu vérifier l'efficacité de la thérapie photodynamique antimicrobienne (TPA) intranasale préopératoire alliée à l'utilisation de lingettes désinfectantes à la chlorhexidine (LDC) pour le corps en prophylaxie des IPO du rachis chez des adultes, calculer les économies réalisées par notre établissement grâce à la stratégie TPA-LDC et recenser ses effets indésirables.

Méthodes : Nous avons procédé à une étude observationnelle de séries chronologiques interrompues échelonnée sur une période de 14 ans chez des adultes (âgés de > 18 ans) soumis à une chirurgie urgente ou non urgente du rachis auprès de 3 cohortes chronologiques spécifiques : avant le déploiement de la stratégie dans notre établissement (2006–2010), puis durant (2011–2014) et après son déploiement (2015–2019). Nous avons appliqué une analyse bivariée non ajustée pour vérifier les changements dans le temps selon la patientèle et les paramètres chirurgicaux, et segmenté la régression pour estimer l'effet de la stratégie TPA-LDC sur les taux d'incidence annuels moyens d'IPO par période. Nous avons utilisé 2 modèles pour estimer le coût de la stratégie TPA-LDC apte à prévenir une IPO additionnelle par année et les économies cumulatives annuelles moyennes générées par la prévention des IPO.

Résultats : Durant la période de l'étude, 13 493 malades (964 par année en moyenne) ont subi une chirurgie urgente ou non urgente du rachis. Entre 2006 et 2011 et entre 2015 et 2019, l'âge moyen, l'indice moyen de comorbidité de Charlson (ICC) et l'indice moyen d'invasivité de la chirurgie du rachis (IICR) sont passés de 48,4 à 58,1 ans, de 1,7 à 2,6 et de 15,4 à 20,5, respectivement ($p < 0,001$). L'analyse non ajustée a confirmé une baisse significative du nombre annuel (de 74,6 à 26,8) et de l'incidence (de 7,98 % à 2,67 %) des IPO associée à la stratégie TPA-LDC ($p < 0,001$). Après ajustement pour tenir compte de l'âge, de l'ICC et de l'IICR moyens, la régression segmentée a montré une réduction absolue du taux d'incidence annuel des IPO de 3,36 % par année ($p < 0,001$). Le coût annuel moyen estimé pour prévenir une IPO additionnelle a été d'environ 1350\$–1650\$, et les économies cumulatives annuelles moyennes estimées ont été de 2 484 856\$–2 495 016\$. Aucun effet indésirable n'a été signalé avec la stratégie TPA-LDC.

Conclusion : L'application préopératoire de la stratégie TPA-LDC est une mesure prophylactique efficace pour les IPO du rachis, et donne lieu à des économies significatives. Compte tenu de sa rapidité d'action, du risque minimal de résistance aux antibiotiques, de son activité à large spectre et du taux d'observance élevé, la décolonisation préopératoire par TPA-LDC devrait être une norme de soins pour l'ensemble de la patientèle soumise à une chirurgie urgente ou non urgente du rachis.

Surgical site infections (SSIs) are devastating postoperative complications associated with substantial morbidity, prolonged length of stay (LOS) and increased health care–related costs.¹ The rate of SSIs after spine surgery is estimated at 1%–16%,^{2,3} with 67.1% of affected patients requiring further surgery.⁴ Multiple patient and surgical risk factors are associated with SSIs after spine surgery.^{2–14} Increased age, greater comorbidity burden, implantation of spinal hardware, greater surgical invasiveness, distant site infection and positive methicillin-resistant *Staphylococcus aureus* (MRSA) cultures are the most consistent.^{2–14} Strategies for preventing SSIs represent an important opportunity to improve patient care and reduce health care–related costs.

Initiatives for reducing the risk of SSIs have been at the forefront of surgical practice and medical literature.^{15–17} Of these, bundled SSI prophylaxis interventions have gained particular interest. Bagga and colleagues¹⁸ reported that the SSI incidence rate was reduced from 3.4% to 1.2% with preoperative chlorhexidine washing and standardized postoperative wound care in a cohort of more than 9000 patients. Featherall and colleagues¹⁹ reported a decrease in the SSI rate from 4.1% to 2.0% with a preoperative bundle that included screening and decolonization for *S. aureus* with nasal mupirocin administration and self-preparation bath with chlorhexidine gluconate (CHG). However, the gold-standard SSI prophylaxis strategy remains unknown, given institutional variability in SSI rates, allocation of resources and cost constraints.

Nasal administration of mupirocin is the decolonization technique for SSI prophylaxis most frequently reported. It has demonstrated efficacy as a single prophylactic measure,²⁰ and Yamada and colleagues²¹ reported a reduction in SSI rates of 3.1% when mupirocin was part of a bundled approach in patients undergoing spine surgery. The accessibility, cost-effectiveness and favourable study outcomes with nasal administration of mupirocin have made it a relatively popular decolonization technique. However, its effectiveness depends on patient compliance and microbio-

logic susceptibility.²² The latter is particularly concerning, given the increasing rate of mupirocin-resistant organisms.^{23,24} Given these limitations, there is a need for an alternative nasal decolonization technique, ideally one without the potential for antimicrobial resistance.

Intranasal photodynamic disinfection therapy (nPDT) is a novel nonantibiotic broad-spectrum antimicrobial technology in which light energy is used to activate a photoactive agent to kill microbial cells in the anterior nares.²² This strategy has been shown to eliminate intranasal carriage of MRSA and methicillin-susceptible *S. aureus*, with sustained decolonization for up to 5 days.²⁵ In 2011, as part of a year-long pilot study at our institution, Bryce and colleagues²⁶ showed that bundled nPDT–CHG was most effective in the spinal surgery population, with an 18-fold reduction in the odds of developing an SSI due to *S. aureus*. The combination of nPDT and CHG decolonization has the potential advantage of broad-spectrum antimicrobial efficacy and a lower risk of antimicrobial resistance.²⁶

Since 2006, we have been prospectively collecting SSI data before, during and after the implementation of nPDT–CHG at our institution. Our primary objective in the present study was to examine the effect of nPDT–CHG on the yearly cumulative incidence rate of SSIs in adult patients undergoing emergent or elective spine surgery. Our secondary objectives were to determine our institutional savings attributable to the use of this strategy and identify adverse events reported with nPDT–CHG.

METHODS

Setting and design

Vancouver General Hospital is a level 1 trauma and quaternary spine referral centre in British Columbia, Canada, serving about 5 million inhabitants. We are the regional centre for complex spine care, including complex adult spinal deformity, trauma and oncology.

This prospective observational interrupted time-series study contains data from Jan. 1, 2006, to Dec. 31, 2019, as part of a quality-improvement initiative for preventing SSIs. Our study population consists of all adult patients (age > 18 yr) undergoing elective or emergent spine surgery for each calendar year.

Preoperative surgical site infection prophylaxis bundle

The preoperative SSI prophylaxis bundle consists of decolonization treatment with CHG-impregnated wipes (Sage Products) to wash the surgical site, axillae and groin within 24 hours preceding surgery, and nPDT (Ondine Biomedical) administered in the preoperative holding area. Patients undergoing elective surgery are given information on the decolonization program in the preadmission clinic or surgeon's office. Patients undergoing emergency surgery are given the information by the admitting surgical service. The nPDT component of the decolonization treatment is administered by trained nurses. This includes applying a photosensitive dye (0.1% methylene blue solution) to the anterior nares for 30 seconds, followed by 2 cycles of 2-minute illumination with a nonthermal red light wavelength of 665 nm.²⁶

Rollout of preoperative surgical site infection prophylaxis bundle

In 2011, a preoperative SSI prophylaxis quality-improvement program was introduced as a pilot study at our institution.²⁶ From Sept. 1, 2011, to Aug. 31, 2012, patients undergoing an elective cardiac, orthopedic, spinal, vascular, thoracic or neurosurgical procedure were recruited to examine the effectiveness of the nPDT-CHG treatment as described above.²⁶

After the success of the pilot study, nPDT-CHG treatment was implemented into full-time practice in 2 phases. Phase I of the rollout (2012–2013) included all elective and emergency surgical procedures performed from 0700 to 1600, Monday to Friday. In phase II of the rollout (2013–2014), the program was gradually expanded to include all elective and emergency surgical procedures performed at any time, any day of the week. By Jan. 1, 2015, all patients undergoing elective or emergent procedures received nPDT-CHG in the preoperative holding area before transfer to the operating room.

Patient and surgical data

We prospectively collected standardized patient and surgical data, including demographic information, Charlson Comorbidity Index (CCI) score, case diagnostic category (deformity, oncology, trauma or degenerative), neurologic status and Spine Surgical Invasiveness Index (SSII) score. The CCI is a validated measure that assesses 19 comorbidities to

produce a weighted score predictive of short- and long-term health outcomes.²⁷ We evaluated surgical invasiveness using the SSII.¹¹ We categorized the SSII scores into mild (< 10), moderate (10–21) and major (> 21) invasiveness.

Outcomes of interest

Our primary outcome of interest was the cumulative change in the annual incidence rate of SSIs after the full-time implementation of preoperative nPDT-CHG, in 2015. We used the US Centers for Disease Control and Prevention National Healthcare Surveillance Network definition of SSI.²⁸ Cases of SSI were confirmed by the attending surgeon and institutional SSI surveillance program (Infection Prevention and Control) using laboratory data, review of the surgical cases list, voluntary reporting, reports from other facilities and review of hospital readmissions with a diagnosis of infection. We collected all cases of confirmed SSI prospectively using our prospective adverse event database, the Spine Adverse Events Severity system, version 2.^{12,13}

Our secondary outcomes of interest were adverse events specific to the administration of nPDT-CHG and the cumulative institutional savings attributable to nPDT-CHG decolonization. We defined adverse events using the Spine Adverse Events Severity system, version 2, a validated, spine-specific, physician-led prospective adverse event database for 14 specific intraoperative and 22 postoperative adverse events^{12,13} (Appendix 1, available at www.canjsurg.ca/lookup/doi/10.1503/cjs.016922/tab-related-content). Adverse events specific to the administration of nPDT-CHG are recorded under the “other” section for each patient.

Cost models

We developed 2 different cost models. The first was an estimate of the nPDT-CHG cost to prevent 1 additional SSI per year. The second was an estimate of cumulative institutional savings after implementation of the SSI prophylaxis program. We derived the base cost of nPDT-CHG per patient of \$45–\$55 from institutional financial reports. The base cost for treating each confirmed case of SSI in our model was \$52 932. We derived this figure from the study by Barnacle and colleagues,²⁹ in which the estimated cost for treating each confirmed case of SSI was NZD\$51 434. We standardized values to 2022 Canadian cost estimates using a 2018 New Zealand dollar to Canadian dollar conversion rate of 0.8973,³⁰ then adjusted for inflation using data from the Bank of Canada.³⁰ We used data from New Zealand to estimate the Canadian equivalent for several reasons. Current Canadian data evaluate the economic impact of pooled medical and surgical adverse events, which dilutes the specific cost of treating spine SSIs.^{31,32} Health economic data across the US literature are variable, with estimates of the cost of SSI treatment per case ranging from US\$16 242¹ to US\$93 741,³³

this range reflects the competing financial interests between federal and private health care institutions and insurance programs.^{33,34} The 2018 cost estimate of Barnacle and colleagues²⁹ represents the most recent health economic data specific for treating spine SSIs, without dilution costs, in a public health care model comparable to that in Canada.^{35,36}

Model 1

We derived the cost of nPDT-CHG to prevent 1 additional SSI per year from the number needed to treat, which we calculated using the adjusted absolute percent change in the annual SSI incidence rate per year from our segmented regression analysis (intercept parameterization region A v. region C). We then multiplied the number needed to treat by the approximate cost of nPDT-CHG per patient, which yielded an average estimate for the total cost of nPDT-CHG to prevent 1 additional SSI per year.

Model 2

We determined the cumulative institutional savings by calculating the differences in the annual SSI cost per year between the prerollout period (2006–2011) and the postrollout period (2015–2019), adjusting for the yearly cost of administering nPDT-CHG from 2015 onward. We calculated the prerollout annual SSI cost per year by multiplying the annual number of confirmed SSIs before rollout by the estimated cost of treating an SSI ($\text{PreC}_T = \text{annual number of SSIs Prerollout} \times \$52\,932$). We calculated the postrollout annual SSI cost per year by multiplying the annual number of confirmed SSIs after rollout of the SSI prophylaxis program by the estimated cost of treating an SSI ($\text{PostC}_T = \text{annual number of SSIs Postrollout} \times \$52\,932$). We calculated the annual cost of administering nPDT-CHG per year after program rollout by multiplying the annual number of surgical cases by the cost of administering nPDT-CHG per patient ($\text{C}_T\text{PDT} = \text{annual surgical cases after rollout per year} \times \$45\text{--}\$55$). Finally, we calculated the annual cumulative institutional savings with adjustment for the cost of administering nPDT-CHG using the formula $\text{InstC}_S = \text{PreC}_T - \text{PostC}_T - \text{C}_T\text{PDT}$.

Statistical analysis

We used descriptive statistics to summarize the patient, surgical and outcome variables of our study population using means, standard deviations (SDs), proportions or rates, as appropriate. We used unadjusted bivariate analysis using the Pearson correlation coefficient to test for temporal changes across individual patient and surgical variables. We used an unadjusted independent 2-sample *t* test to test the difference in the mean number and incidence of SSIs per year between the pre- and postrollout periods.

We developed a segmented regression model to determine the effect of the nPDT-CHG program on the yearly SSI incidence rate, adjusting for independent predictors of

SSI.³⁷ The SSI incidence rate was recorded every calendar year from 2006 to 2019. We developed the segmented regression model with 2 joinpoints: 2011 (last year of the prerollout period) and 2014 (last year of the rollout period). The model was constrained to meet at each joinpoint, while allowing for different year slopes to be fitted within regions A (prerollout), B (rollout) and C (postrollout). The model included the covariates mean age, mean CCI score and mean SSII score.

We parameterized the model in several ways to numerically determine the effect of the nPDT-CHG program on the yearly SSI incidence rate. First, we parameterized the model to calculate the slope (β) representing the adjusted annual SSI incidence rate per year for each region. We then parameterized the model to compare the difference in slopes between regions, representing the adjusted change in the annual SSI incidence rate per year between the prerollout, rollout and postrollout periods. Finally, we parameterized the model by comparing the difference in intercepts (α) between regions, representing the adjusted absolute percent change in the annual SSI incidence rate per year between the prerollout, rollout and postrollout periods. We confirmed model equivalence numerically between the various parameterizations by checking the other effects in the models. A 95% confidence interval (CI) was provided with all effect estimates from the segmented regression analysis. We performed the analyses using SAS version 9.4 (SAS Institute).

Ethics approval

This study was approved by the Vancouver Coastal Health Regional Ethics Committee and the University of British Columbia Clinical Research Ethics Board (H10-03441) as a quality-improvement project. The study was conducted following the principles of the Helsinki Declaration. The requirement for informed consent was waived because of the anonymous nature of the data.

RESULTS

Patient demographic characteristics and case distribution

Over the study period, 13 493 patients underwent elective or emergent spine surgery at our institution, with 4 670 cases before program rollout, 3 741 during the rollout period and 5 082 after program rollout. There was a mean of 964 (SD 73) surgical cases per calendar year (Table 1). From 2006 to 2019, the mean age increased by a decade, from 48.4 to 58.1 years ($r = 0.98$, $p < 0.001$). The proportion of patients older than 75 years of age and older than 85 years of age increased from 45.3% to 51.2% and from 6.5% to 9.8%, respectively ($r = 0.93$, $p < 0.001$) (Figure 1). The mean CCI score and mean SSII score increased from 1.7 to 2.6 ($r = 0.97$, $p < 0.001$) and from 15.4 to 20.5 ($r = 0.88$, $p < 0.001$), respectively (Figure 2).

Table 1. Patient demographic characteristics by year

Year	No. of patients <i>n</i> = 13 493	Mean age ± SD, yr	% of patients > 75 yr of age	% of patients > 85 yr of age	% of patients female	Mean CCI score ± SD	% of patients with neurodeficit
2006	978	48.4 ± 12.4	45.3	6.5	54.1	1.7 ± 0.9	16.4
2007	875	49.8 ± 14.7	46.0	6.7	53.4	1.8 ± 1.1	14.2
2008	938	51.5 ± 11.3	46.8	5.9	48.8	1.7 ± 0.8	16.9
2009	984	50.6 ± 13.2	45.3	6.3	55.3	1.9 ± 0.8	13.3
2010	895	52.8 ± 14.9	44.9	6.4	52.6	1.8 ± 1.0	14.0
2011	925	53.5 ± 13.4	47.4	6.9	51.9	1.9 ± 1.1	12.5
2012	903	53.7 ± 12.8	48.1	7.2	51.2	2.0 ± 0.9	10.7
2013	921	53.4 ± 13.4	47.7	8.1	54.5	2.2 ± 1.1	14.9
2014	992	54.1 ± 12.1	48.8	7.8	48.9	2.3 ± 1.0	15.7
2015	965	55.8 ± 12.7	49.2	8.6	52.7	2.3 ± 0.8	15.3
2016	916	55.5 ± 14.2	50.4	9.5	52.1	2.5 ± 1.2	13.0
2017	978	56.4 ± 14.8	51.1	9.4	51.8	2.6 ± 1.1	15.3
2018	1105	57.2 ± 13.6	52.6	8.9	53.5	2.5 ± 1.2	15.8
2019	1118	58.1 ± 13.7	51.2	9.8	54.2	2.6 ± 0.9	16.1

CCI = Charlson Comorbidity Index; SD = standard deviation.

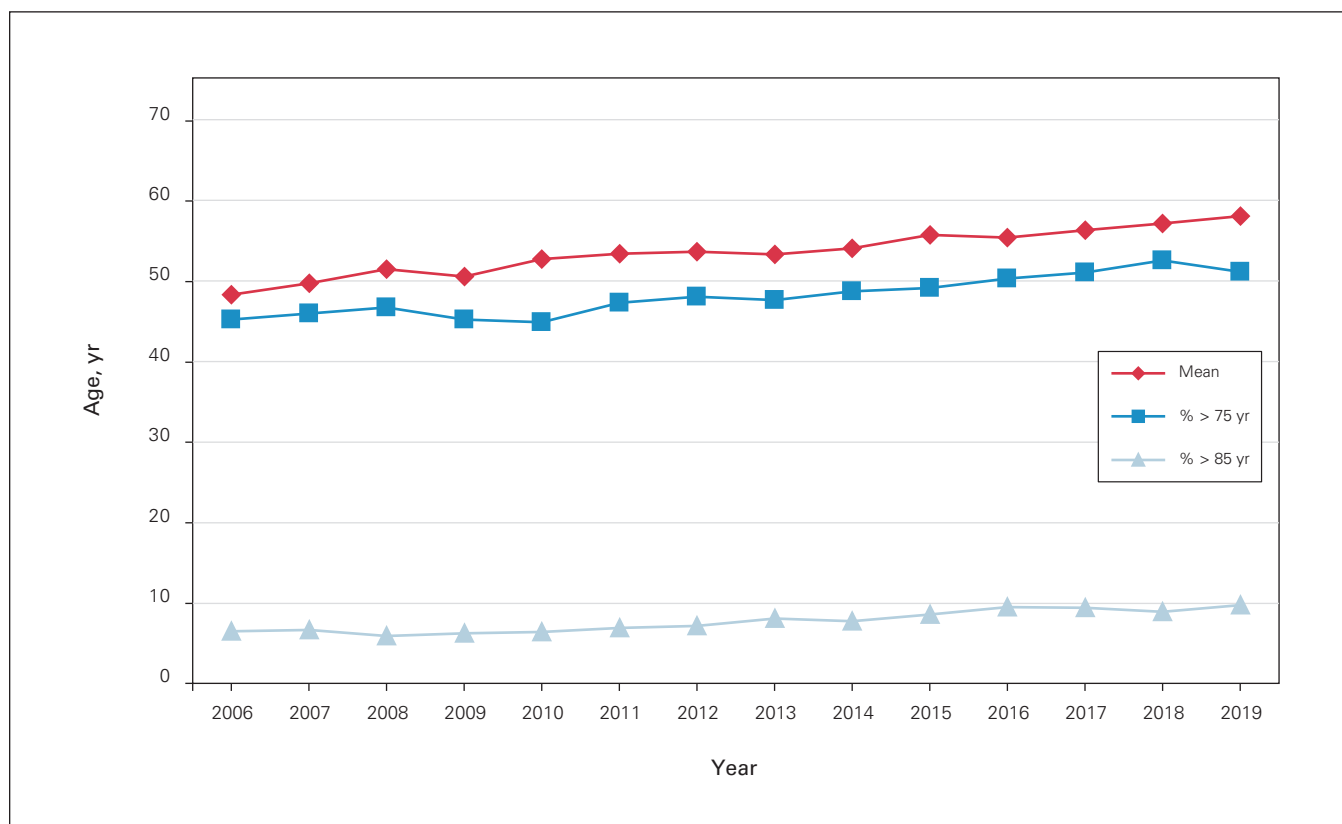


Fig. 1. Patient age distribution over the study period.

The proportion of SSII scores greater than 21 increased from 52.5% to 61.0% ($r = 0.82$), and the proportion of SSII scores less than 10 decreased from 14.0% to 8.0% ($r = -0.90$) ($p < 0.001$). The proportion of oncology cases increased from 8.0% to 11.0% ($r = 0.66$, $p = 0.011$), and the proportion of

deformity and trauma cases decreased from 16.0% to 13.5% ($r = -0.64$, $p = 0.01$) and from 23.0% to 19.0% ($r = -0.587$, $p = 0.03$), respectively (Table 2 and Figure 3). There were no changes in the proportion of degenerative surgical cases during the study period ($r = -0.33$, $p = 0.2$).

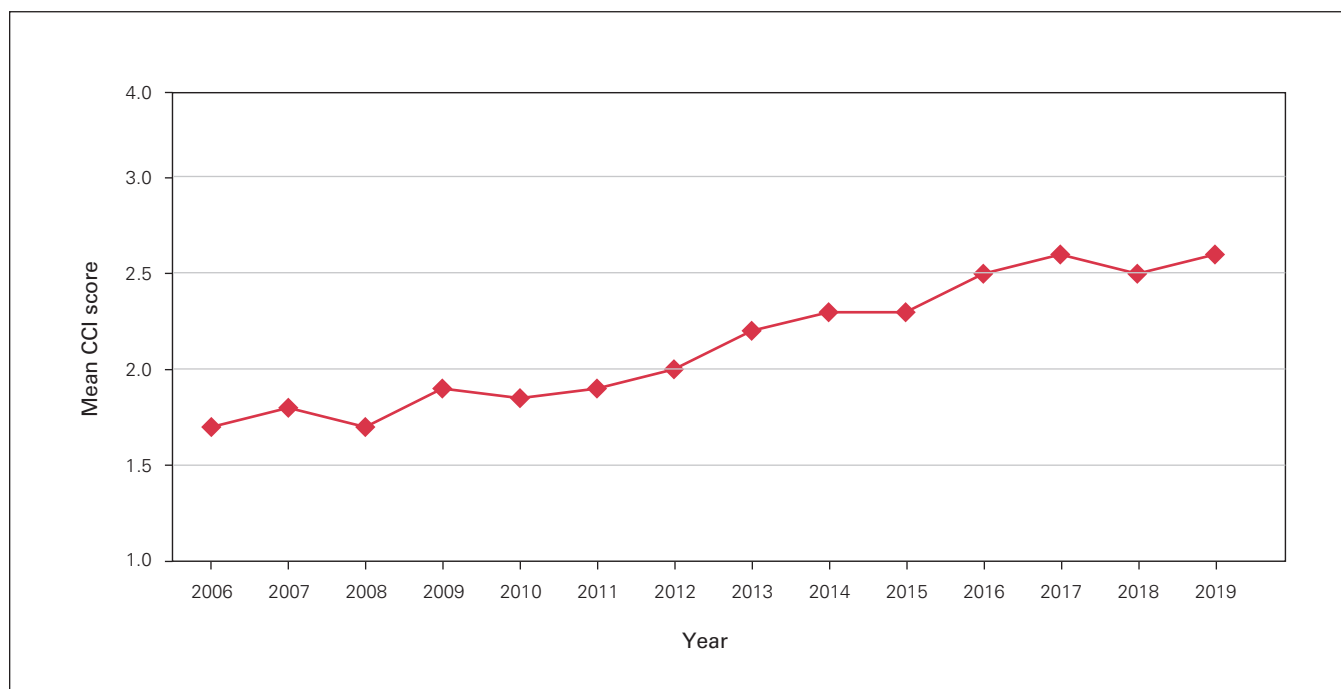


Fig. 2. Mean Charlson Comorbidity Index (CCI) score over the study period.

Table 2. Surgical demographic characteristics by year

Year	Diagnostic category; no. (%*) of patients 13493						Mean SSII score \pm SD	% of patients with SSII score < 10	% of patients with SSII score > 21
	Total <i>n</i> = 13 493	Trauma <i>n</i> = 3300	Oncology <i>n</i> = 1395	Deformity <i>n</i> = 1997	Degenerative <i>n</i> = 4435	Other <i>n</i> = 2366			
2006	978	225 (23.0)	78 (8.0)	156 (16.0)	342 (35.0)	177 (18.1)	15.4 \pm 5.6	13.9	52.5
2007	875	245 (28.0)	66 (7.5)	123 (14.1)	297 (33.9)	144 (16.5)	17.6 \pm 6.3	13.3	52.0
2008	938	253 (27.0)	80 (8.5)	169 (18.0)	281 (30.0)	155 (16.5)	18.0 \pm 5.7	13.4	51.9
2009	984	236 (24.0)	103 (10.5)	157 (16.0)	339 (34.4)	149 (15.1)	18.7 \pm 6.1	12.5	56.5
2010	895	219 (24.5)	98 (10.9)	134 (15.0)	300 (33.5)	144 (16.1)	18.8 \pm 6.8	11.4	61.3
2011	925	245 (26.5)	97 (10.5)	134 (14.5)	315 (34.0)	134 (14.5)	19.1 \pm 6.8	12.3	60.0
2012	903	235 (26.0)	113 (12.5)	131 (14.5)	316 (35.0)	108 (12.0)	18.9 \pm 7.1	8.9	57.8
2013	921	244 (26.5)	88 (9.6)	147 (16.0)	299 (32.5)	143 (15.5)	18.4 \pm 8.8	10.5	60.7
2014	992	253 (25.5)	104 (10.5)	159 (16.0)	298 (30.0)	178 (17.9)	19.1 \pm 8.3	8.4	59.2
2015	965	241 (25.0)	111 (11.5)	140 (14.5)	318 (33.0)	155 (16.1)	20.6 \pm 9.2	8.5	58.5
2016	916	229 (25.0)	110 (12.0)	120 (13.1)	289 (31.6)	168 (18.3)	20.3 \pm 9.7	9.3	61.4
2017	978	220 (22.5)	108 (11.0)	127 (13.0)	318 (32.5)	205 (21.0)	19.9 \pm 9.4	9.9	60.5
2018	1105	243 (22.0)	116 (10.5)	149 (13.5)	343 (31.0)	254 (23.0)	20.5 \pm 9.9	7.3	61.8
2019	1118	212 (19.0)	123 (11.0)	151 (13.5)	380 (34.0)	252 (22.5)	20.5 \pm 9.1	8.4	61.3

SD = standard deviation; SSII = Spine Surgery Invasiveness Index.
*By year total.

Pearson correlation analysis identified significant associations between mean age and mean CCI score ($r = 0.91$, $p < 0.001$), mean age and proportion of oncology cases ($r = 0.69$, $p = 0.001$), and proportion of oncology cases and mean SSII score ($r = 0.75$, $p = 0.002$). Overall, over the study period, there was a statistically significant increase in mean patient age of a decade; in patient comorbidity burden, by almost 50%; in surgical invasiveness, by 20%; and in the proportion of surgical oncology cases, by 40%.

Cumulative surgical site infection incidence rate

Before program rollout, the mean number of confirmed SSIs per year was 74.6 (SD 5.18), with a maximum SSI incidence rate of 8.21% and a minimum incidence rate of 7.71% (Table 3 and Figure 4, region A). During the rollout period, the mean number of confirmed SSIs per year was 47.75 (SD 15.78), with a maximum SSI incidence rate of 7.35% and a minimum incidence rate of 3.12% (Figure 4, region B). After rollout, the mean number of confirmed

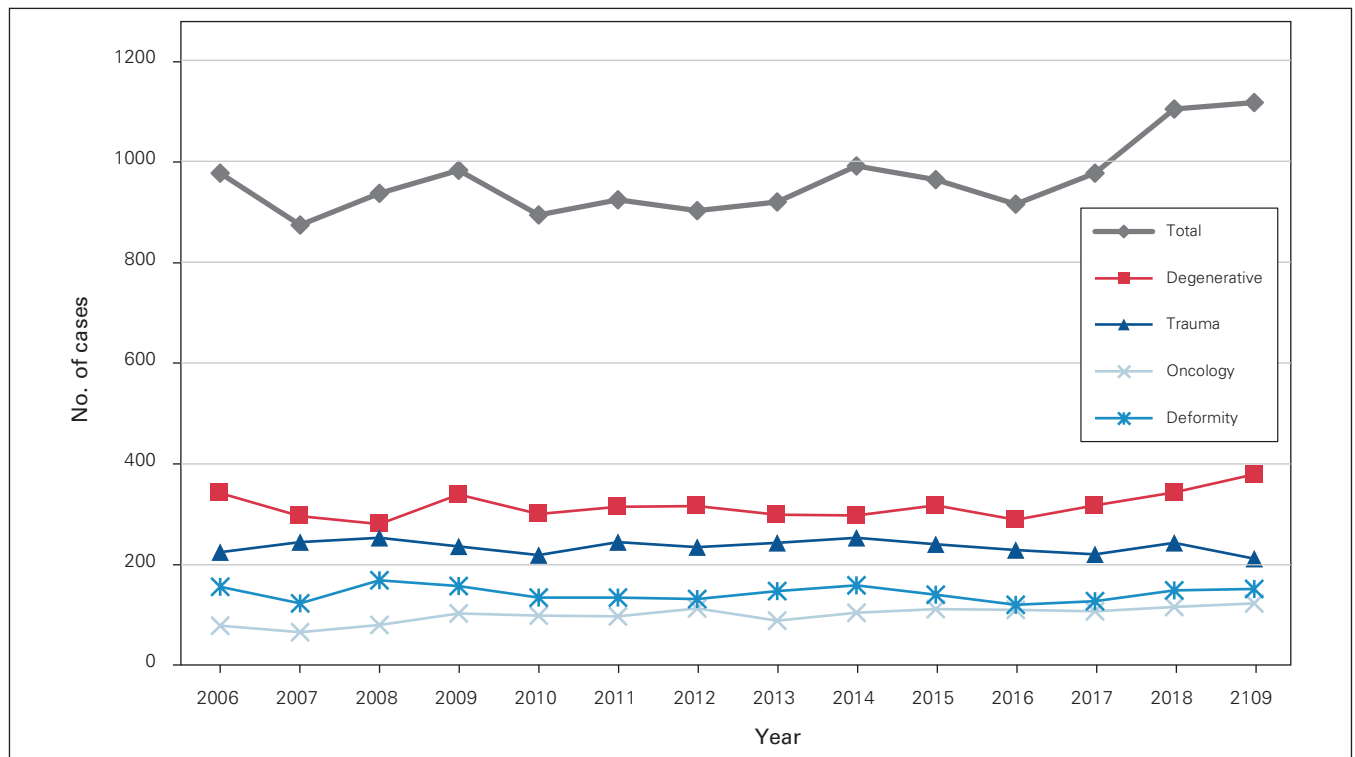


Fig. 3. Distribution of cases by diagnostic category over the study period.

Year	No. (%) of SSIs
2006	79 (8.1)
2007	69 (7.9)
2008	77 (8.2)
2009	79 (8.0)
2010	69 (7.7)
2011	68 (7.4)
2012	51 (5.6)
2013	41 (4.4)
2014	31 (3.1)
2015	34 (3.5)
2016	26 (2.8)
2017	25 (2.6)
2018	25 (2.3)
2019	24 (2.2)

SSI = surgical site infection.

SSIs per year was 26.8 (SD 4.09), with a maximum SSI incidence rate of 3.52% and a minimum incidence rate of 2.15% (Figure 4, region C). Unadjusted analysis confirmed a significant reduction in the mean number and mean incidence of SSIs per year between the pre- and postrollout periods ($p < 0.001$).

Segmented regression analysis with slope parameterization identified that, after adjustment for mean age, mean

CCI score and mean SSII score, program rollout (region B) was associated with a statistically significant reduction in the annual SSI incidence rate of 1.32% (95% CI -1.95 to -0.69) per year (Table 4). Parameterization of slope differences between regions confirmed that the greatest percent change in the annual SSI incidence rate occurred only during rollout, compared to the pre- or postrollout period ($p < 0.001$). Intercept parameterization between the prerollout period (region A) versus the postrollout period (region C) showed that, after adjustment for mean age, mean SSII score and mean CCI score, program rollout was associated with a statistically significant absolute percent reduction in the annual SSI incidence rate of 3.36% (95% CI -4.67 to -2.04) per year. None of the adjustment variables were statistically significant except for mean SSII score, which had borderline significance ($p = 0.06$).

Adverse events

No adverse events associated with nPDT-CHG were reported or observed during the study.

Cost of intervention and cumulative institutional savings

With model 1, after adjustment for mean age, mean CCI score and mean SSII score, the estimated annual cost associated with nPDT-CHG to prevent 1 additional spine SSI per year during program rollout was about \$1350–\$1650.

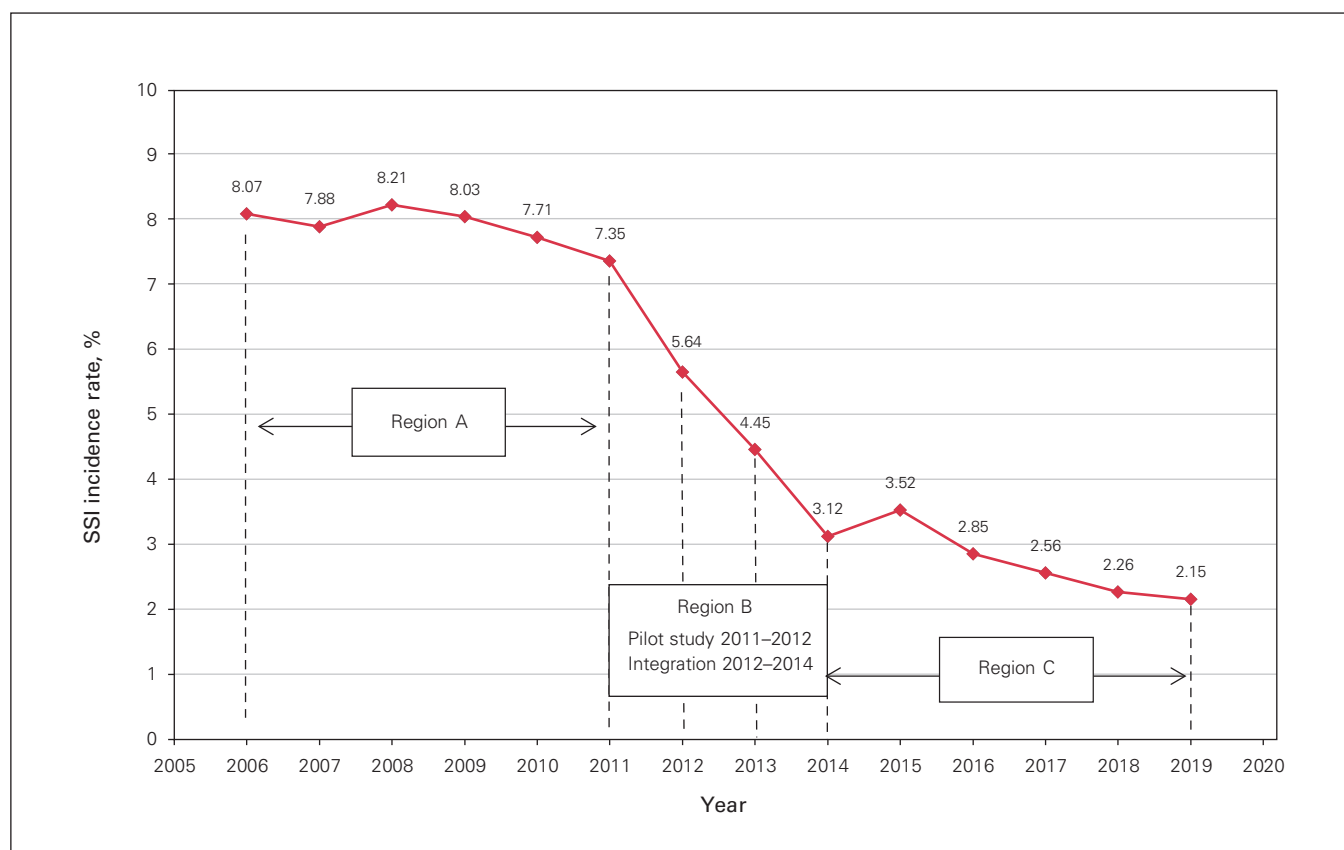


Fig. 4. Cumulative surgical site infection (SSI) incidence rate over the study period. Region A: before rollout of the SSI prophylaxis program; region B: program rollout; region C: after program rollout.

Table 4. Segmented regression analysis*

Variable	Estimate (95% CI)
Baseline adjustment variables	
Mean age	0.01 (−0.47 to 0.49)
Mean CCI score	−0.52 (−4.19 to 3.15)
Mean SSII score	0.28 (−0.02 to 0.58)
% change in SSI incidence rate per year	
Region A	−0.32 (−0.82 to 0.19)
Region B	−1.32 (−1.95 to −0.69)
Region C	−0.27 (−0.81 to 0.26)
% change in SSI incidence rate per year between regions†	
Region B v. region A	−1.00 (−1.40 to −0.61)
Region C v. region A	0.04 (−0.14 to 0.23)
Region B v. region C	1.05 (0.71 to 1.38)
Absolute % change in SSI incidence rate per year between regions	
Region A v. region B	5.01 (3.05 to 6.97)
Region A v. region C	−3.36 (−4.67 to −2.04)
Region B v. region C	−8.37 (−11.03 to −5.70)

CI = confidence interval; CCI = Charlson Comorbidity Index; SSI = surgical site infection; SSII = Spine Surgery Invasiveness Index.
 *Segmented regression model adjusted for increase in mean age, mean CCI score and mean SSII score over the study period.
 †Region A: prerollout period (Jan. 1, 2006, to Dec. 31, 2010); region B: rollout period (Jan. 1, 2011, to Dec. 31, 2014); region C, postrollout period (Jan. 1, 2015, to Dec. 31, 2019).

With model 2, the estimated annual prerollout SSI treatment cost was \$3 969 900 per year, and the estimated mean postrollout treatment cost was \$1 429 164. The annual cost of administering nPDT-CHG after rollout was \$45 720–\$55 880. Therefore, the estimated annual cumulative institutional savings attributable to the use of nPDT-CHG was \$2 484 856–\$2 495 016 per year.

DISCUSSION

Our study showed that the preoperative use of nPDT-CHG was associated with a significant reduction in the annual SSI incidence rate of 5.31% per year. After adjusting for age, surgical invasiveness and comorbidity burden, we observed a significant reduction in the annual SSI incidence rate of 1.32% per year during nPDT-CHG rollout and an absolute reduction in the annual yearly SSI incidence rate of 3.36% per year. The use of nPDT-CHG also offered estimated cumulative institutional savings of \$2 484 856–\$2 495 016 annually. In addition, no adverse events were reported, which suggests that nPDT is a safe technology.

The cumulative reduction in the incidence of SSIs occurred during significant changes in our study population over time. Concurrent with the aging population, we observed a significant increase in patient age of a decade,

a 50% increase in mean CCI score, a 20% increase in mean SSII score and a 40% increase in the proportion of oncology surgery cases performed each year. The CCI score is one of the strongest independent risk factors for SSI after spine surgery, with a reported hazard ratio of 2.48,³⁸ and older age almost doubles the odds of developing an SSI.^{39,40} Moreover, we observed an increase in overall mean SSII scores and proportion of SSII scores greater than 21, which may reflect the yearly increase in oncology cases. A significant dose–response relation exists between SSII score and infection, with a score greater than 21 increasing the relative risk of SSI by 3.15 times.⁹ Patients undergoing spine surgery for oncologic conditions also experience a greater incidence of SSI, with reported rates as high as 4.4%⁴¹ and 9.5%.⁴² Despite significant increases in all these risk factors in our population, bundled preoperative nPDT–CHG remained an effective intervention.

There are limited studies investigating the effect of nPDT–CHG. Our results are consistent with a meta-analysis of the use of similar preoperative bundled SSI prevention strategies.⁴³ Yamada and colleagues²¹ observed a reduction in the SSI rate from 3.8% to 0.7% and an absolute risk reduction of 84% with nasally administered mupirocin and CHG skin decolonization. Similarly, Schweizer and colleagues⁴⁴ reported a risk reduction of 42% with a bundled SSI prevention strategy in complex SSI due to *S. aureus* after cardiac, hip or knee surgery. We observed that, after full implementation of the nPDT–CHG program, the slope representing the annual SSI incidence rate plateaued, with a loss of statistical significance. The plateau observed in the postrollout period likely reflects the maximal preventive effect of nPDT–CHG in patients undergoing spinal surgery. The insignificant difference in adjusted slope between the pre- and postrollout periods also suggests that, after full implementation of the program, there was a return to the baseline variability in yearly SSI incidence. This residual baseline variability, saturated by the maximal preventive effect of nPDT–CHG, represents the baseline SSI risk attributable to the medical and surgical complexity of our patient population devoid of colonization. These findings further validate the effectiveness of nPDT–CHG in reducing the risk of SSI associated with nasal and skin colonization in a high-risk population.

Since full implementation of the nPDT–CHG program, we have implemented other changes in surgical practice at our institution to prevent SSIs, such as intrawound administration of vancomycin powder, wound drains, negative pressure wound dressings and insertion of silver-coated Foley catheters. These factors were not accounted for in the present study, as they occurred only after rollout and their use is highly variable depending on surgeon preference, case complexity and patient factors. Furthermore, our analysis identified that the only significant change in yearly SSI incidence rate occurred during nPDT–CHG rollout, which validates the effectiveness of

this SSI prophylaxis strategy. This is in keeping with the finding of Bryce and colleagues²⁶ that nPDT–CHG decolonization resulted in greater reduction in SSI incidence in patients undergoing elective cardiac, orthopedic, spinal, vascular, thoracic or neurosurgical procedures than in a historical control group and a concurrent control group, reducing the odds of SSI due to *S. aureus* by 18 fold.

The effectiveness of nasally administered mupirocin depends on patient compliance and antibiotic susceptibility.^{23,26} Nicolas and colleagues⁴⁵ reported that a therapeutic mupirocin concentration was found in the nasal secretions of only 22 of 41 patients who reported good compliance with nasal self-administration of mupirocin. Hetem and colleagues²³ reported a rate of mupirocin resistance in coagulase-negative *Staphylococcus* of 96% after nasal decolonization with mupirocin. In comparison, preoperative nPDT has been found to have excellent patient compliance (94%), ease of use without interrupting nursing workflow and an average administration time of 10 minutes.²⁶ Bundled nPDT and CHG skin decolonization also offers the theoretical advantage of broad-spectrum antimicrobial efficacy, rapid action and low risk of development of antimicrobial resistance.²⁶

Our cost analysis identified that the estimated cost of nPDT–CHG to prevent 1 additional SSI per year was \$1350–\$1650 per year. Our estimated annual institutional cumulative savings attributable to nPDT–CHG use was \$2 484 856–\$2 495 016 per year, adjusted for the ongoing annual cost of administering nPDT–CHG to all patients after program rollout. Interestingly, our cumulative institutional savings were substantially greater than those previously reported. Stambough and colleagues⁴⁶ estimated a net savings of US\$717 205.59 within 25 months after implementing a decolonization program that reduced the incidence of SSI from 0.8% to 0.2% in patients undergoing elective total hip and knee replacement. Rennert-May and colleagues⁴⁷ reported a cost saving of \$153 per person and 16 SSIs avoided annually with nasally administered mupirocin and CHG skin decolonization in more than 8000 hip or knee replacement procedures. The institutional savings observed in our study reflects the effectiveness of nPDT–CHG in a high-risk surgical population and the avoided substantial costs of treating spine SSIs.²⁹ Accordingly, bundled nPDT–CHG may be a financially viable long-term solution, given its affordability and clinical effectiveness.

We observed no additional adverse events specific to nPDT–CHG use in our prospective adverse event database. Bryce and colleagues²⁶ reported 7 cases of transient pharyngeal irritation, likely related to trickling of the methylene blue into the back of the throat, but no other complications. Two of the 7 patients were referred to otorhinolaryngology for nasopharyngoscopy, and no tissue reaction was observed on examination. Those authors did not report any cases of altered taste or smell related to the nasal treatment.

Limitations

Our findings must be interpreted within the context of the study design. Its prospective nature, large sample, long follow-up period and direct comparison of pre- versus post-intervention cohorts, together with the fact that rigorous SSI surveillance continued unaltered after full-time implementation of the nPDT-CHG program, lend strength to our findings. In addition, the interrupted time-series design and use of segmented regression analysis are validated methodologies for evaluating the effectiveness of large-scale health interventions implemented at specific times and for testing causal hypotheses about the intervention.⁴⁸⁻⁵⁰

Although no additional interventions were introduced during nPDT-CHG rollout, our study did not account for variabilities and changes in other practice standards. The increase in SSI awareness and prevention has introduced changes in nursing care, medical education and generalized ward management. Although these factors cannot be accounted for, it remains that the significant reduction in SSIs occurred only during nPDT-CHG rollout.

Surveillance data are constrained by the lack of individualized data. Accordingly, we could not account for the number of patients who received nPDT-CHG from 2011 to 2014. However, as nPDT-CHG was a defined intervention with specific independent time points, the difference in SSI prevention cost between the pre- and postrollout periods provides a reasonably accurate estimate of our institutional savings attributable to the use of nPDT-CHG. Comparing defined time points also eliminates any assumptions made during nPDT-CHG rollout that may skew the cost estimates. Likewise, by incorporating the ongoing annual cost of administering nPDT-CHG to all patients after rollout, our model reflects a more accurate estimate of our institutional savings. Last, the scope of SSI prevention has a far greater impact than cost savings alone: it also substantially reduces the physical, emotional and medical burden on patients and families.

Finally, our study did not investigate whether there were differences or changes in the etiologic organism or antimicrobial resistance patterns in patients who developed an SSI. These are clinically important questions to address, as they provide insight into whether preoperative nPDT-CHG is associated with changes in the medical and surgical management of patients who develop an SSI postoperatively. Future studies are needed to determine whether preoperative nPDT-CHG decolonization is associated with changes in etiologic SSI organism(s), antimicrobial resistance patterns, length of antibiotic treatment or differences in surgical management⁵¹ such as the number of surgical washouts and/or hardware revisions.

CONCLUSION

Preoperative bundled nPDT-CHG is a clinically effective strategy for reducing the incidence of SSIs after emergent or

elective spine surgery. It is an affordable intervention and is associated with significant institutional savings for every SSI prevented in this high-risk population. Given its rapid action, minimal risk of antimicrobial resistance, broad-spectrum activity and high compliance rate, preoperative bundled nPDT-CHG decolonization should be the standard of care for all patients undergoing emergent or elective spine surgery.

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References

- Blumberg TJ, Woelber E, Bellabarba C, et al. Predictors of increased cost and length of stay in the treatment of postoperative spine surgical site infection. *Spine J* 2018;18:300-6.
- Yao R, Tan T, Tee JW, et al. Prophylaxis of surgical site infection in adult spine surgery: a systematic review. *J Clin Neurosci* 2018;52:5-25.
- Olsen MA, Nepple JJ, Riew KD, et al. Risk factors for surgical site infection following orthopaedic spinal operations. *J Bone Joint Surg Am* 2008;90:62-9.
- Haleem A, Chiang HY, Vodola R, et al. Risk factors for surgical site infections following adult spine operations. *Infect Control Hosp Epidemiol* 2016;37:1458-67.
- Meng F, Cao J, Meng X. Risk factors for surgical site infections following spinal surgery. *J Clin Neurosci* 2015;22:1862-6.
- Pull ter Gunne AF, Cohen DB. Incidence, prevalence, and analysis of risk factors for surgical site infection following adult spinal surgery. *Spine* 2009;34:1422-8.
- Fei Q, Li J, Lin J, et al. Risk factors for surgical site infection after spinal surgery: a meta-analysis. *World Neurosurg* 2016;95:507-15.
- Lieber B, Han B, Strom RG, et al. Preoperative predictors of spinal infection within the National Surgical Quality Inpatient Database. *World Neurosurg* 2016;89:517-24.
- Cizik AM, Lee MJ, Martin BI, et al. Using the spine surgical invasiveness index to identify risk of surgical site infection: a multivariate analysis. *J Bone Joint Surg Am* 2012;94:335-42.

10. Liu JM, Deng HL, Chen XY, et al. Risk factors for surgical site infection after posterior lumbar spinal surgery. *Spine* 2018;43:732-7.
11. Mirza SK, Deyo RA, Heagerty PJ, et al. Towards standardized measurement of adverse events in spine surgery: conceptual model and pilot evaluation. *BMC Musculoskelet Disord* 2006;7:53.
12. Rampersaud YR, Neary MA, White K. Spine Adverse Events Severity system: content validation and interobserver reliability assessment. *Spine* 2010;35:790-5.
13. Street JT, Lenehan BJ, DiPaola CP, et al. Morbidity and mortality of major adult spinal surgery. A prospective cohort analysis of 942 consecutive patients. *Spine J* 2012;12:22-34.
14. DiPaola CP, Saravanja DD, Boriani L, et al. Postoperative infection treatment score for the spine (PITSS): construction and validation of a predictive model to define need for single versus multiple irrigation and debridement for spinal surgical site infection. *Spine J* 2012;12:218-30.
15. Puffer RC, Murphy M, Maloney P, et al. Increased total anesthetic time leads to higher rates of surgical site infections in spinal fusions. *Spine* 2017;42:E687-90.
16. Blood AG, Sandoval MF, Burger E, et al. Risk and protective factors associated with surgical infections among spine patients. *Surg Infect (Larchmt)* 2017;18:234-49.
17. Nasser R, Kosty JA, Shah S, et al. Risk factors and prevention of surgical site infections following spinal procedures. *Global Spine J* 2018;8:44s-48s.
18. Bagga RS, Shetty AP, Sharma V, et al. Does preventive care bundle have an impact on surgical site infections following spine surgery? An analysis of 9607 patients. *Spine Deform* 2020;8:677-84.
19. Featherall J, Miller JA, Bennett EE, et al. Implementation of an infection prevention bundle to reduce surgical site infections and cost following spine surgery. *JAMA Surg* 2016;151:988-90.
20. Bode LG, Kluytmans JA, Wertheim HF, et al. Preventing surgical-site infections in nasal carriers of *Staphylococcus aureus*. *N Engl J Med* 2010;362:9-17.
21. Yamada K, Abe H, Higashikawa A, et al. Evidence-based care bundles for preventing surgical site infections in spinal instrumentation surgery. *Spine* 2018;43:1765-73.
22. Septimus EJ, Schweizer ML. Decolonization in prevention of health care-associated infections. *Clin Microbiol Rev* 2016;29:201-22.
23. Hetem DJ, Bootsma MC, Bonten MJ. Prevention of surgical site infections: Decontamination with mupirocin based on preoperative screening for *Staphylococcus aureus* carriers or universal decontamination? *Nephrol Dial Transplant* 2016;62:631-6.
24. Septimus EJ. Nasal decolonization: What antimicrobials are most effective prior to surgery? *Am J Infect Control* 2019;47S:A53-7.
25. Street C, Pedigo L, Gibbs A, et al. Antimicrobial photodynamic therapy for the decolonization of methicillin-resistant *Staphylococcus aureus* from the anterior nares. In: Kessel DH, editor. *Photodynamic therapy: back to the future*. 12th World Congress of the International Photodynamic Association. Vol. 7380 of *Proceedings* series. Bellingham (WA): SPIE; 2009.
26. Bryce E, Wong T, Forrester L, et al. Nasal photodisinfection and chlorhexidine wipes decrease surgical site infections: a historical control study and propensity analysis. *J Hosp Infect* 2014;88:89-95.
27. Charlson ME, Pompei P, Ales KL, et al. A new method of classifying prognostic comorbidity in longitudinal studies: development and validation. *J Chronic Dis* 1987;40:373-83.
28. Surgical site infection (SSI). Atlanta: Centers for Disease Control and Prevention. Available: <https://www.cdc.gov/HAI/ssi/ssi.html> (accessed 2023 May 7).
29. Barnacle J, Wilson D, Little C, et al. Excess cost and inpatient stay of treating deep spinal surgical site infections. *N Z Med J* 2018;131:27-34.
30. Inflation calculator. Ottawa: Bank of Canada; 2022. Available: <https://www.bankofcanada.ca/rates/related/inflation-calculator/> (accessed 2022 Oct. 30).
31. Hellsten EK, Hanbidge MA, Manos AN, et al. An economic evaluation of perioperative adverse events associated with spinal surgery. *Spine J* 2013;13:44-53.
32. Rampersaud YR, Power JD, Perruccio AV, et al. Healthcare utilization and costs for spinal conditions in Ontario, Canada — opportunities for funding high-value care: a retrospective cohort study. *Spine J* 2020;20:874-81.
33. Edmiston CE, Leaper DJ, Chitnis AS, et al. Risk and economic burden of surgical site infection following spinal fusion in adults. *Infect Control Hosp Epidemiol* 2023;44:88-95.
34. O'Neill J, O'Neill D. Health status, health care and inequality: Canada vs. the U.S. *Forum Health Econ Policy* 2008;10:1094.
35. Interactive tool OECD. International comparisons — peer countries, Canada. Available: <https://www.cihi.ca/en/oeed-interactive-tool-international-comparisons-peer-countries-canada> (accessed 2023 May 7).
36. Thet N, Thet Hsu Hnin N, Mya N. Comparison of the international health care systems through the consideration of population health and performance indicators in Canada, Australia and New Zealand: a systematic literature review. *Int J Sci Res Publ* 2021;11:199-205.
37. Bernal JL, Cummins S, Gasparrini A. Interrupted time series regression for the evaluation of public health interventions: a tutorial. *Int J Epidemiol* 2017;46:348-55.
38. Kurtz SM, Lau E, Ong KL, et al. Infection risk for primary and revision instrumented lumbar spine fusion in the Medicare population. *J Neurosurg Spine* 2012;17:342-7.
39. Kim J, Kim TH. Risk factors for postoperative deep infection after instrumented spinal fusion surgeries for degenerative spinal disease: a nationwide cohort study of 194,036 patients. *J Clin Med* 2022;11:778.
40. Nagashima H, Nanjo Y, Tanida A, et al. Clinical features of spinal infection in individuals older than eighty years. *Int Orthop* 2012;36:1229-34.
41. Shillingford JN, Laratta JL, Reddy H, et al. Postoperative surgical site infection after spine surgery: an update from the Scoliosis Research Society (SRS) Morbidity and Mortality Database. *Spine Deform* 2018;6:634-43.
42. Omeis IA, Dhir M, Sciubba DM, et al. Postoperative surgical site infections in patients undergoing spinal tumor surgery: incidence and risk factors. *Spine* 2011;36:1410-9.
43. Wolfhagen N, Boldingh QJJ, Boermeester MA, et al. Perioperative care bundles for the prevention of surgical-site infections: meta-analysis. *Br J Surg* 2022;109:933-42.
44. Schweizer ML, Chiang HY, Septimus E, et al. Association of a bundled intervention with surgical site infections among patients undergoing cardiac, hip, or knee surgery. *JAMA* 2015;313:2162-71.
45. Nicolas R, Carricajo A, Morel J, et al. Evaluation of effectiveness and compliance with the mupirocin nasal ointment part of *Staphylococcus aureus* decolonization in real life using UPLC-MS/MS mupirocin quantification. *J Antimicrob Chemother* 2020;75:1623-30.
46. Stambough JB, Nam D, Warren DK, et al. Decreased hospital costs and surgical site infection incidence with a universal decolonization protocol in primary total joint arthroplasty. *J Arthroplasty* 2017;32:728-734.e721.
47. Rennert-May E, Conly J, Smith S, et al. A cost-effectiveness analysis of mupirocin and chlorhexidine gluconate for *Staphylococcus aureus* decolonization prior to hip and knee arthroplasty in Alberta, Canada compared to standard of care. *Antimicrob Resist Infect Control* 2019;8:113.
48. Schaffer AL, Dobbins TA, Pearson SA. Interrupted time series analysis using autoregressive integrated moving average (ARIMA) models: a guide for evaluating large-scale health interventions. *BMC Med Res Methodol* 2021;21:58.
49. Taljaard M, McKenzie JE, Ramsay CR, et al. The use of segmented regression in analysing interrupted time series studies: an example in pre-hospital ambulance care. *Implement Sci* 2014;9:77.
50. Kontopantelis E, Doran T, Springate DA, et al. Regression based quasi-experimental approach when randomisation is not an option: interrupted time series analysis. *BMJ* 2015;350:h2750.
51. Dowdell J, Brochin R, Kim J, et al. Postoperative spine infection: diagnosis and management. *Global Spine J* 2018;8:37s-43s.